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Defaults versus framing: Revisiting Default Effect and Framing Effect with a replication and extension of Johnson and Goldstein (2003) and Johnson, Bellman, and Lohse (2002)

Subramanya Prasad Chandrashekar^{*1}, Nadia Adelina^{*2}, Shiyuan Zeng^{*2}, Yan Ying Esther

Chiu^{*2}, Grace Yat Sum Leung^{*2}, Paul Henne³, Boley Cheng², and Gilad Feldman²

¹Department of Psychology, Norwegian University of Science and Technology (NTNU)

²Department of Psychology, University of Hong Kong, Hong Kong SAR

³Department of Philosophy, Lake Forest College

*Joint first authors

People tend to stick with a default option instead of switching to another option. For instance, Johnson and Goldstein (2003) found a default effect in an organ donation scenario: if organ donation is the default option, people are more inclined to consent to it. Johnson et al. (2002) found a similar default effect in health-survey scenarios: if receiving more information about your health is the default, people are more inclined to consent to it. Much of the highly cited, impactful work on these default effects, however, has not been replicated in well-powered samples. In two well-powered samples (N = 1920), we conducted a close replication of the default effect in Johnson and Goldstein (2003) and in Johnson et al. (2002). We successfully replicated Johnson and Goldstein (2003). In an extension of the original findings, we also show that default effects are unaffected by the permanence of these selections. We, however, failed to replicate the findings of Johnson et al. (2002)'s study; we did not find evidence for a default effect. We did, however, find a framing effect: participants who read a positively-framed scenario consented to receive health-related information at a higher rate than participants who read a negatively framed scenario. We also conducted a conceptual replication of Johnson et al. (2002) that was based on an organ-donation scenario, but this attempt failed to find a default effect. Our results suggest that default effects depend on framing and context. Materials, data, and code are available on: https://osf.io/8wd2b/.

Keywords: action framing effect; default effect; organ donation; nudge; replication; choice; judgment and decision making

Suppose that people receive a health survey after a doctor's appointment in order to see if they would like to receive health updates from their doctors. If the option to participate is preselected, people would probably not change their response—instead sticking with the default option and participating in the service. This is an example of the default effect: given a default option, people stick with it rather than changing (Johnson and Goldstein, 2003; Johnson et al., 1993).

The framing of the options may also affect people's choices. In this example, people would be more inclined to select an option if it is framed positively, as in answering "Yes" to "I will participate," as opposed to negatively, as in selecting "No" to "I will not participate." This is an example of a framing effect: people consent to participate at a higher rate when a choice is positively framed than when it is negatively framed (Johnson and

Goldstein, 2003).

Default effects and framing effects have been very influential across many academic disciplines and in public policy (Araña and León, 2013; Evans et al., 2011; Johnson and Goldstein, 2003; Mintz and Redd, 2003; Tversky and Kahneman, 1981). The use of default effects is a well-known effective example of leveraging behavioral insights to influence people or to nudge them toward specific socially desirable choices. Governments and public policy organizations worldwide have set-up Nudge Units that implemented interventions using default effects to encourage desired behavior encouraging organ donations and pension savings (Halpern, 2015).

There is, however, some evidence for an overestimation of the size of nudge effects. For instance, DellaVigna and Linos (2022) recently found that there were larger effect sizes for nudge interventions reported in published literature than those reported by Nudge Units in the United States. This finding suggests that selective reporting may lead to inflated meta-analytic effect sizes (Kvarven et al., 2020). Moreover, in some cases, nudge effects did not replicate with larger samples (Bohner and Schlüter, 2014;Kettle et al., 2017; Kristal et al., 2020).

Given these recent findings, there is reason to investigate default effects and framing effects. Despite a substantial number of experimental studies on default effects, for instance, very few of these employed preregistered analysis plans using well-powered samples (Szaszi et al., 2018). Together, these may lead to misplaced optimism about easy-to-implement nudging interventions, while much more complex solutions involving structural reforms have been ignored (Schmidt and Engelen, 2020). As such, researchers have called for more preregistered replications using well-powered samples (Ferguson and Heene, 2012; Franco et al., 2014).

In the current research, we sought to revisit and reassess classics on default and framing effects by embarking on preregistered high-power replications and extensions of two impactful studies on default effects: Johnson and Goldstein (2003) and Johnson et al. (2002). The first study by Johnson and Goldstein (2003) was an early demonstration of default effects. The study found that people were more likely to register as organ donors when the default option was to register. Johnson et al. (2002) contrasted default effects against framing effects and found that default effects prevailed, and that framing did not change the participants' tendency to select the default over alternatives. We investigated these foundational studies.

Default Effect

Early demonstrations of default effect were in the context of auto-insurance choices made in New Jersey and Pennsylvania, when each state had a different policy regarding the right to sue for damages in auto accidents (Johnson et al., 1993). New Jersev residents had low insurance premiums yet could acquire an additional right to sue at an additional cost. Pennsylvanian residents by default had the right to sue, but they could opt out of this right and pay a lower insurance premium. For instance, Johnson et al. (1993) found that 75% of Pennsylvania auto-insurance consumers paid the higher premium and retained their right to sue. In comparison, only 20% of New Jersey auto-insurance consumers actively chose to pay the additional premium and obtain the right to sue. Researchers have since found support for the default effect in a variety of contexts related to health, retirement saving, organ donation. sustainability, insurance coverage, electricity consumption, charitable giving, and many other decision-making domains (Abadie and Gay, 2006; Benartzi and Thaler, 1999;Cronqvist and Thaler, 2004; Ebeling, 2013; Jachimowicz et al., 2019; Madrian and Shea, 2001; Shealy and Klotz, 2015)¹. While a few studies failed to support default effects (Abhyankar et al., 2014; Everett et al., 2015; Keller et al., 2011; Reiter et al., 2012), a recent meta-analysis noted substantial variations in the efficacy of the default effects (Jachimowicz et al., 2019); for instance, defaults in consumer domains were more effective, while defaults in environmental domains were less effective (Jachimowicz et al., 2019).

Framing Effects

People's decisions are also influenced by the way a decision scenario is framed-whether by using different wordings, settings, or situations (Brewer and Kramer, 1986; De Martino et al., 2006; Fagley and Miller, 1987; Gamliel and Kreiner, 2013; Huber et al., 1987; Kramer, 1989; Kühberger, 1998; Levin and Gaeth, 1988; Piñon and Gambara, 2005; Puto, 1987; Rothman and Salovey, 1997). Johnson et al. (2002) tested the action framing effects of a decision by manipulating whether participants were asked to select (positive frame) or reject (negative frame) an option. Participation rates in the positively framed condition were higher than the negatively framed condition. In this case, the positive or negative framing greatly influenced people decisions. The findings are consistent with the view that positive dimensions of a choice are weighted more when selecting an option whereas the negative dimensions are weighted more when rejecting an option (Shafir et al., 1993).

Present research

We selected Johnson and Goldstein (2003) and Johnson et al. (2002) as our replication targets for three reasons: each is foundational, has been highly influential in academia (Kahneman, 2003; Kruglanski and Gigerenzer, 2011; Weber and Johnson, 2009), and has been highly influential in practice for policy making.

Johnson and Goldstein (2003)'s work was the first to demonstrate the use of defaults in an organ donation scenario, and at the time of writing this article the paper has been cited more than 2000 times. In the original study, the experimenters varied whether the donor or non-donor status was the default option. Organ donation rates were higher when the default option was

¹Although not directly relevant to the current study, researchers have offered a variety of explanations for default effects (e.g., Brown and Krishna, 2004; Huh et al., 2014; Johnson and Goldstein, 2003; McKenzie et al., 2006)

to donate (82%) than when the default option was to not donate (42%). These findings have influenced public policy decisions; Argentina (Nacion, 2005), Uruguay (Trujillo, 2013), Chile (Zúñiga-Fajuri, 2015), England (English et al., 2019), and Wales (Griffiths, 2013; Madden et al., 2020) have adopted default organ-donor status policies. Organ donation statistics from the Organization for Economic Cooperation and Development (OECD) countries show that, on average, organ donation rates are higher in countries where the default option is to donate (Opt-Out system) than in countries where the default option was not to donate (Opt-In system) (Li and Nikolka, 2016).

To the best of our knowledge, Johnson et al. (2002) were the first to investigate the interaction of framing of action (we refer to this framing effect here as an action framing effect)² and default effects in people's decisions. In the original study, the researchers asked participants whether they would like to be notified about future health surveys after they completed an online health questionnaire (Johnson et al., 2002). The experimenters varied whether the default selection was to receive these future notifications, not to receive these future notifications, or neither. They also varied whether the options were framed positively ("Notify Me") or negatively ("Do Not Notify Me"). Consistent with the default effect, participants were more inclined to be notified when participation was the default. Although the framing manipulation was not significant as a predictor of participants' decision to receive these future notifications, the pattern of responses showed that participants in the positive framing conditions consented to receive health-related information at a higher rate than participants in the negative framing conditions (Johnson et al., 2002).

We embarked on direct well-powered replications of these two classic findings with two primary goals. First, we sought to revisit and reexamine the robustness of the basic default effect reported in the well-known organ donation decision scenario by Johnson and Goldstein (2003). Second, to build on these findings we sought to also contrast default and framing effects, replicating and extending the design used in Johnson et al. (2002).

Effect Sizes in target articles

The chosen target studies did not report effect sizes. We reanalyzed the data and conducted logistics regression analysis to calculate odds-ratios with a 95% confidence interval for the regression coefficients as a measure of effect size. The effect sizes of the original studies are summarized in Table 9 (for detailed results, see Table S7 and Table S8 in the supplementary materials).

Extensions

In addition to the direct replications, we also performed two extensions. First, we investigated whether the permanence of the decision affects default effects. In particular, half the study participants were told that their organ donation-related decision was valid for three years, and the other half of participants were not provided with any additional information about the permanence of their decision. We based our extension on van Dalen and Henkens (2014) who found that organ donation rates were higher when the option was temporary and would have to be renewed than when the default option was to donate. Based on these results, we investigated the presumed permanence (temporary vs. permanent) consent in Johnson and Goldstein (2003) scenario. In line with previous work, we predicted a higher organ donation participation rate when the choices were framed as temporary (i.e., the decision can be revised in five years) rather than permanent. Second, we added a conceptual replication of Johnson et al. (2002). We applied their experimental design involving framing and default effects to the organ-donation scenario in Johnson and Goldstein (2003). This replication was meant to further test the generalizability of their findings regarding the interaction of default and framing effects.

Method

Process

We crowdsourced the replication and extension effort using two teams of two authors. Both teams were supervised by two other experienced authors. Each team worked independently to conduct their own in-depth analysis of the chosen target articles and wrote detailed pre-registrations aiming for a very close replication and adding additional extensions. Data collection was then conducted separately for each team using a different sample. Thus, the two data collections tested two independent extensions: the effect of choice permanence (Sample 1) and the conceptual replication of Experimental 2 of Johnson et al. (2002) (Sample 2).

Pre-registrations and open data/code

In both data collections, we first preregistered the experiment on the Open Science Framework (OSF) and data collection was launched after registration. Preregistration, disclosures, power analyses, and all materials are available in the Supplementary Materials. These, together with datasets and analysis code, were

²For an *action framing effect*, the presentation of a scenario varies in the framing of the action (e.g., to select vs. to reject).

made available on the OSF at https://osf.io/8wd2b/. All measures, manipulations, and exclusions for this investigation are reported, and data collection was completed before analyses. Pre-registrations are available on the OSF: Group A - https://osf.io/mhwbe/, Group B - https://osf.io/j4rpc/.

Participants and power analysis

The present investigation includes two simultaneously collected data samples. For both the samples, we recruited participants from the United States via CloudResearch platform running on Amazon Mechanical Turk. Participants could participate in only one of these.

Power analyses across Group A and Group B suggested a sample size of 232. However, we note inconsistencies in the power analysis details reported as part of the pre-registrations across Group A and B. Rectified power analysis based on the original study's results indicated that a total sample of 156 participants was sufficient to obtain 95% power (at $\alpha = .05$) to detect the smallest effect reported among the original studies (*OR* = 1.86). Please refer to the supplementary material for more details on the power analysis.

Since our replication study also involved additional extension hypotheses across two samples, we recruited 1004 and 1007 participants across two replication two teams, respectively. Additionally, a post-hoc power sensitivity analysis at an aimed sample size of 2000 participants is found to achieve a power of 96.93% power (at = .05) to detect a small effect size (i.e., OR = 1.50). We, therefore, combined the two samples for the data analysis, which amounted to a total of 2011 participants. Following the preregistered exclusion criteria, we excluded 91 participants based on English proficiency, self-reported seriousness, knowledge of the hypothesis, survey completion, and place of residence (see supplementary material for details). Data were analyzed from the remaining 1920 participants (N1 = 954; N2 = 966; $M_{age} = 38,SD = 11.85;52\%$ female).

Materials and Procedure

The procedure involved two parts. In the first part, participants read about an organ donation scenario from Johnson and Goldstein (2003). In the second part, participants responded to the scenario from Experiment 2 of Johnson et al. (2002). After completing both parts of the survey, participants provided their demographic information, and they were debriefed at the end of the study. We provide a comparison of the target article sample and the replication samples in Table S2. Participants in Sample 1 were part of a choice permanence extension. In this regard, Sample 1 participants in the

first of the experiment were randomly assigned to one of two between-participants conditions: the direct replication of Johnson and Goldstein (2003) or the temporary organ donation extension condition.

Part 1: Organ Donation

In part 1, participants were randomly assigned to 1 of 3 between-participants conditions (defaults: Opt-In vs. Opt-Out vs. No-Default). For example, participants in the "Opt-Out" condition read:

"Imagine that you have just moved to a new state and are currently filling out the required online registration forms when you are asked to indicate your organ donor status. The default in this state is that you ARE automatically enrolled to be an organ donor. You are given the choice of whether to confirm or to change this status. Please select an option."

After reading the passage, participants had to choose either "Yes - I want to be an organ donor" or "No - I do not want to be an organ donor." In the Opt-Out condition, the "Yes" option was pre-selected. Table 1 documents the format of the display of choices across experimental conditions. So, participants who consented to organ donation just had to click "Next" at the bottom of the page, whereas participants who did not wish to be an organ donor had to click the option "No" before clicking "Next." In the Opt-In condition, the "No" option was pre-selected. So, participants who consented to organ donation had to click the option "Yes" before proceeding, whereas participants who did not wish to be an organ donor just had to click "Next." In the No-Default condition, participants read:

Assume you moved to a new state, therefore, you need to select enrollment as an organ donor. Please choose your preferred organ donor status:

Participants in this No-Default condition saw the same binary response options without a pre-selected default. So, they had to actively select "Yes" or "No" before clicking "Next" to proceed. After completing part 1, participants moved on to part 2.

Part 2: Survey Subscription

In part 2, participants were randomly assigned to 1 of 6 conditions in a 2 (framing: Positive vs. Negative) \times 3 (default option: Opt-In vs. Opt-Out vs. No-Default) between-participants design (see Table 2 for details). At the beginning of Part 2, participants read the following instruction:

"Typically, regardless of your organ donor decision, the state online systems ask you to answer a number of health questions. Please answer the following. All the data will be kept completely confidential."

Study stimuli for the direct replication of Johnson and Goldstein (2003)

[Introduction for participants in Opt-Out/Opt-in Conditions]:

Imagine that you have just moved to a new state and are currently filling out the required online registration forms when you are asked to indicate your organ donor status. The default in this state is that you ARE automatically enrolled to be an organ donor. You are given the choice of whether to confirm or to change this status. Please select an option

[Opt-out]:

Assume you moved to a new state in which the default is that you are an organ donor, you are therefore by default enrolled as an organ donor. Please choose your preferred organ donor status:

- Yes-I want to be an organ donor
- \bigcirc No- I do not want to be an organ donor

[Opt-in]:

Assume you moved to a new state in which the default is that you are not an organ donor, you are therefore by default not enrolled as an organ donor. Please choose your preferred organ donor status:

- \bigcirc Yes-I want to be an organ donor
- No- I do not want to be an organ donor

[No-default]:

Assume you moved to a new state, therefore, you need to select enrollment as an organ donor. Please choose your preferred organ donor status:

- \bigcirc Yes-I want to be an organ donor
- \bigcirc No- I do not want to be an organ donor

Participants then answered four generic questions on their health in general (for details, see Table S4 supplementary section). Participants then read:

"You are almost at the end of the survey. Thank you for taking part. Would you be interested in being notified about other policy/health-related surveys? (If yes, we will contact you through MTurk using your MTurk worker ID)"

Participants answered by selecting "Yes" or "No." Each condition had a positive ("Notify me about more health surveys.") or negative ("Do NOT notify me about more health surveys.") framing. In positively framed Opt-Out conditions, the 'yes' response was pre-selected. In positively framed Opt-In conditions, the 'No' response was pre-selected. In negatively framed Opt-Out conditions, the 'No' response was pre-selected. In negatively framed Opt-in conditions, the 'yes' response was preselected.

Extensions

Extension 1: The effect of choice permanence.

Participants in Sample 1 were part of the choicepermanence extension. As such, participants in Sample 1 were randomly assigned to one of two betweenparticipants conditions (temporary or permanent). Participants assigned to the temporary conditions took the same survey as those in the permanent conditions—only they received the following additional instruction at the beginning of part 1 of the study:

"Please note: Your organ donor authorization, if granted, would be for 3 years only, meaning that after 3 years you will be asked to reconfirm your organ donor decision."

Participants in the permanent conditions had no additional instructions.

Extension 2: Conceptual replication of Experimental 2 of Johnson et al. (2002).

All the participants in Sample 2 took part in a different extension. Immediately after completing Part 1 of the study but just before Part 2, participants read the following instructions (see Table 3 for details):

"Would you like to receive further information about organ donation through MTurk? If you indicate your approval, we'll contact you through MTurk using your worker ID with further information about organ donation."

These participants were randomly assigned to 1 of 6 conditions in a 2 (framing: Positive vs. Negative) times 3 (default option: Opt-Out vs. Opt-In vs. No-Default) between-participants design (for details, see Table S6 in the supplementary section). After reading the above instruction, participants selected "Yes" or "No" to a question asking for consent to receiving further information on organ donation. Each of the default conditions involved either a positive ("Send me more information about organ donation") or negative ("Do NOT send me more information about organ donation") framing. The responses were pre-selected in the Opt-In and Opt-Out default conditions mirroring the experimental design of Experiment 2 of Johnson et al. (2002). In positively framed Opt-Out conditions, the 'yes' response was preselected. In positively framed Opt-In conditions, the 'No' response was pre-selected. In negatively framed Opt-Out conditions, the 'no' response was pre-selected.

Study stimuli for the direct replication of Johnson et al. (2002)

[Introduction]:

Typically, regardless of your organ donor decision, the state online systems ask you to answer a number of health questions. Please answer the following. All the data will be kept completely confidential.

You are almost at the end of the survey. Thank you for taking part. Would you be interested in being notified about other policy/health-related surveys? (If yes, we will contact you through MTurk using your MTurk worker ID)

[Positive frame, Opt-out]:

Notify me about more health surveys.

• Yes

 \bigcirc No

[Positive frame, Opt-in]: Notify me about more health surveys.

○ Yes

• No

[Positive frame, No-default]:

Notify me about more health surveys.

 \bigcirc Yes

 \bigcirc No

[Negative frame, Opt-out]: Notify me about more health surveys.

○ Yes

• No

[Negative frame, Opt-in]: Do NOT notify me about more health surveys.

• Yes

 \bigcirc No

[Negative frame, No-default]: Notify me about more health surveys.

○ Yes

 \bigcirc No

In negatively framed Opt-In conditions, the 'yes' response was pre-selected.

Data Transformations

Both Part 1 and Part 2 of the replication study collected the participants' responses in a binary format (Yes/No). In the organ donation scenario, we coded the answer "Yes" (i.e., consenting to donate organs) as '1'. We coded the response "No" (i.e., dissenting to donate organs) as '0'. In Part 2, the responses that indicate the consent for participation were coded as "1," whereas non-participation was coded as "0". Coding of the answers for the choice permanence extension and the conceptual replication of Johnson et al. (2002) followed the same response coding procedure as that of Part 1 and Part 2, respectively. Consistent with previous literature, we define default effects as the difference in participation rates between the Opt-Out condition versus that in the Opt-In condition. We then calculated the odds-ratio for the regression coefficients as a measure of effect size with a 95% confidence interval.

Analysis

Data were analyzed using R. Data were fit to logistic regression models using the *glm* function (with "binomial("logit")" as the family). In Part 1, we report the results from 2-sample test for equality of proportions (comparing the participation rate across different defaults). We analyzed the data for Part 2 with a 2 *times* 3 binomial logistic regression, with framing (Positive vs. Negative) and default (Opt-In vs. Opt-Out vs. NoDefault) and interaction terms of framing and defaults as predictors of the respondent's decision to participate (1 = Yes; 0 = No). For testing the effect of choice permanence on organ donation rate, we conducted chi-square tests comparing the participation rate across temporary and permanent conditions. In the second extension, we conducted the same analysis as for part 2 of the study.

Replication evaluation

We evaluated findings in our replication effects using the criteria set by LeBel et al. (2019) (see Table S15 and Figure S2 in the supplementary material). Table 4 provides a classification of the replications using the LeBel et al. (2019) criteria. We summarize the present replications as very close.

Results

We provide a summary of the findings in Table 9. We present complete descriptive statistics across the two samples in Table 5 (also see Table S10 in the supplementary materials).

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Study stimuli for the on conceptual replication of Johnson et al. (2002)

[Introduction]:

Typically, regardless of your organ donor decision, the state online systems ask you to answer a number of health questions. Please answer the following. All the data will be kept completely confidential.

You are almost at the end of the survey. Thank you for taking part. Would you be interested in being notified about other policy/health-related surveys? (If yes, we will contact you through MTurk using your MTurk worker ID)

[Positive frame, Opt-out]:

Send me more information about organ donation.

• Yes

○ No

[Positive frame, Opt-in]:

Send me more information about organ donation.

- Yes
- No

[Positive frame, No-default]:

Send me more information about organ donation.

○ Yes

○ No

[Negative frame, Opt-out]:

Do NOT send me more information about organ donation.

 \bigcirc Yes

• No

[Negative frame, Opt-in]:

Do NOT send me more information about organ donation.

- Yes
- \bigcirc No

[Negative frame, No-default]:

Do NOT send me more information about organ donation.

○ Yes

○ No

Table 4

Table 4. Classification of replications based on LeBel et al.(2019)

Design facet	Replication study
IV operationalization	Same
DV operationalization	Same
IV stimuli	Same
DV stimuli	Same
Procedural details	Similar
Physical settings	Different
Contextual variables	Different
Replication classification	Very close replication

Part 1: Replication of Johnson and Goldstein (2003)

Consistent with the original study, participants in the Opt-Out condition consented to organ donation at a higher rate (73.5%) than participants in the Opt-In condition (62.5%) (Chi-squared test: $\chi^2(1) = 12.96$, p < .001, Odds ratio (OR) = 1.67, 95% CI [1.27, 2.19] (see Figure 1). This result was consistent across both samples (See Table S11 for complete results). Also, participants in the No-Default condition consented to organ donation at a higher rate (69.7%) than participants in the Opt-In condition (62.5%) ($\chi^2(1) = 5.31$, p = .021, OR = 1.38, 95% CI [1.06, 1.80]) with slight deviations between the two samples (See Table 6 for the results based on logistic regression.

Part 2: Replication of Johnson, Bellman, and Lohse (2002)

We present the results of the regression analysis in Table 7 (Figure 2), and descriptive statistics in Table S9 in the supplementary section.

Default effects

We failed to find support for differences in rates of consent to receive health-related information between the Opt-Out condition (60.5%) and the Opt-In condition (61.1%) (b = -.29, p = .095, OR = 0.75, 95% CI [0.53, 1.05]); that is, we found no support for a default effect. This result was consistent across both samples (See Table S11 for complete results). Participants in the No-Default (59.8%) condition, moreover, consented to receive health-related information at a lower rate than participants in Opt-In (61.1%) condition (b = -0.41, p = .021, OR = 0.67, 95% CI [0.47, 0.94]).

Descriptive table of the participation rates.

Replication Study	Experimental Conditions	,	Combi	ined replication sample
Replication Study	Experimental Conditions	,	n	Participation rate
	Opt-in default		488	62.5%
Replication of Experiment 1 from Johnson & Goldstein (2003)	Opt-out default		476	73.5%
	No-default (no default)		482	69.7%
	Opt in default	Positive Framing	324	88.6%
	Opt-in default	Negative Framing	324	33.6%
Penlication of Experiment 2 from Johnson Bellman & Johse (2002)	Opt out default	Positive Framing	321	93.1%
Replication of Experiment 2 from Johnson, bennan & Lonse (2002)	Opt-out default	Negative Framing	319	27.6%
	No default (no default)	Positive Framing	320	93.4%
		Negative Framing	312	25.3%

Note. N = 1920.

Table 6

Summary of the replication results of Part 1 (Johnson and Goldstein (2003)) based on logistic regression analysis

Predictor	Estimate	SE	Z	р	OR [95% CI]
Intercept	0.51	0.09	5.46	<.001	1.67 [1.39, 2.00]
Default: No-Default – Opt-In	0.32	0.14	2.37	0.018	1.38 [1.06, 1.80]
Default: Opt-Out – Opt-In	0.51	0.14	3.65	<.001	1.67 [1.27, 2.19]

Note. Estimates represent the odds of the dependent variable = "1" vs. "0".



Figure 1

Results of direct replications of Johnson and Goldstein (2003). Percentage of participants consenting to organ donation by condition across both samples. Note.* p < .05, ** p < .01, *** p < .001

Framing effects

Participants in the positive framing condition consented to receive health-related information (91.7%) at a higher rate than participants in the negative framing condition (28.9%) (b= 2.74, p < .001, OR = 15.61, 95% CI [10.31, 23.63]).

Exploratory: Default effects as a function of frames

We proceeded to conduct additional exploratory (not preregistered) analyses examining the interaction between framing and defaults. We found support for an interaction (see Table 7). We considered two sets of interactions: 1) (Positive – Negative) *times* (No-default – Opt-In); 2) (Positive – Negative) *times* (Opt-Out – Opt-In).

For the (Positive – Negative) × (No-default – Opt-In) interaction, we looked at the consent rates between the No-Default and Opt-In conditions across the positive and negative frame (b = 1.01, p = .003, OR = 2.76, 95% CI [1.43, 5.32]). Within positive framing conditions, participants in the No-Default condition (93.4%) consented to receive health-related information at a higher rate than participants in the Opt-In condition (88.6%) (see Table S12 and Table S13 in the supplementary materials). The pattern of results was in the opposite direction in the negative framing conditions: participants in the No-Default condition (25.3%) consented to receive health-related information at a lower rate than participants in the Opt-In condition (33.6%).

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The results were similar for the (Positive – Negative) *times* (Opt-Out – Opt-In) interaction (b = 0.85, p = .010, OR = 2.35, 95% CI [1.23, 4.49]). Within positive framing conditions, participants in the Opt-Out condition (93.1%) consented to receive health-related information at a higher rate than participants in the Opt-In condition (88.6%). The pattern of results was in the opposite direction within the negative framing conditions: participants in the Opt-Out condition (27.6%) consented to receive health-related information at a lower rate than participants in the Opt-In condition (33.6%).

Extensions

Extension 1: Temporary vs. Permanent Choice

We found no evidence that organ donation rates varied between the temporary (70.3%) and the permanent conditions (65.8%) (χ^2 (1, N = 954) = 1.94, p = .163; d = 0.11, 95% CI [-0.04, 0.26]; Figure 3). Additionally, we looked at differences within the No-Default condition, and again failed to find evidence for differences (temporary: 74.7%; permanent: 67.3%; χ^2 (1, N = 320) = 1.78, p = .182, d = 0.20, 95% C.I. [-0.07, 0.47]).

Extension 2: Conceptual replication of Experimental 2 of Johnson et al. (2002)

We summarized the regression analysis in Table 8 and Figure 4, and descriptive statistics are provided in Table S14 in the supplementary.

Participants in the Opt-Out condition consented to receive organ donation-related information (50.2%) at a higher rate than participants in Opt-In condition (37.3%) (b = 0.59; p = .012, OR = 1.81, 95% CI [1.14, 2.87]). We found no evidence that consent rate varied between participants in the No-Default condition (46.6%) and participants in the Opt-In condition (37.3%) (b = 0.39; P = .092, OR = 1.48, 95% CI [0.94, 2.32]). Participants in the positive framing condition consented to receive organ donation-related information (24.9%) at a lower rate than participants in negative framing condition (64.6%) (b = 1.82; p < .001, OR = 0.16, 95% CI [0.10, 0.27]). We found no evidence that results of defaults on participation rate vary as a function of frame (see Table 8).

Summary of replication findings

We replicated the default effects from Johnson and Goldstein (2003). In Part 1 of our study, participants consented to donate their organs at a higher rate when they had to opt out relative to when they had to opt in. We, however, failed to replicate the default effects in Johnson et al. (2002). In Part 2 of our study, we found no evidence that consent to be notified about health-related surveys varied between the opt-out and opt-in conditions. Furthermore, we found that people in positively framed scenarios consented to receive health-related information at a higher rate than participants in negatively framed scenarios. This result deviated from the findings of Johnson et al. (2002) that showed no framing effects.

We followed LeBel et al.'s (2019) framework for the evaluation of our replication using three factors: (a) whether a signal was detected (i.e., the confidence interval for the replication Effect Size (ES) excludes one), (b) consistency of the replication ES with the original study's ES, and (c) precision of the replication's ES estimate (see Figure S2 in the supplementary material). We summarized our evaluations of the replications' findings based on LeBel et al.'s (2019) replication evaluation framework in Table 9 (see Figure 5).

Extensions: Summary of findings

In the first extension, we predicted that people would be more inclined to become donors when consent to organ donation is temporary. We found no evidence that consent varied between the temporary and permanent conditions.

In the second extension, we conducted a conceptual replication of Experiment 2 of Johnson et al. (2002) using the scenario from Part 1 in which participants consented to receive additional information about organ donation. We found support for the default effect: participants who had to opt out consented at higher rates than those who had to opt in. Deviating from the original study where they found no support for framing effects, we found that people in positively framed scenarios consented to receive health-related information at a lower rate than participants in negatively framed scenarios. Framing effects in our extension is opposite to those found in our direct replication of the original study scenario (Johnson et al. 2002).

Summary of findings of Johnson et al. (2002) across original, direct replication, and conceptual replication studies

The findings across direct and conceptual replication of Johnson et al. (2002) provide mixed support to the original study's assertion. We summarize the comparison of the findings in Table 10. Both direct and conceptual replication failed to find differences in consent rates between the No-Default condition and the Opt-In condition. Only the conceptual replication found that consent rates were higher in Opt-out condition than in the Opt-In condition. Regarding the framing effects, we 10

Table 7

Summary of the replication results of Part 1 (Johnson Goldstein, 2003) based on logistic regression analysis

Prodictor			Model 1				Model 2	
Fredictor	Estimate	Ζ	р	OR [95% CI]	Estimate	Ζ	р	OR [95% CI]
Intercept	-0.84 (0.11)	-7.65	<.001	0.43 [0.35, 0.53]	-0.68 (0.12)	-5.78	<.001	0.51 [0.40, 0.64]
Framing: Positive – Negative	3.31 (0.14)	24.14	<.001	27.27 [20.97, 35.88]	2.73 (0.21)	12.96	<.001	15.30 [10.23, 23.40]
Default: No-Default – Opt-In	-0.12 (0.15)	-0.81	.417	0.89 [0.66, 1.19]	-0.40 (0.18)	-2.29	.021	0.66 [0.47, 0.94]
Default: Opt-Out – Opt-In	-0.05 (0.15)	-0.35	.724	0.95 [0.71, 1.27]	-0.29 (0.17)	-1.66	.096	0.75 [0.54, 1.05]
Interaction term:								
(Positive – Negative) \times (No-Default–Opt-In)					1.00 (0.33)	3.02	.003	2.74 [1.43, 5.35]
(Positive – Negative) × (Opt-Out –Opt-In)					0.85 (0.33)	2.57	.010	2.33 [1.23, 4.49]

Note. Estimates represent the odds of the dependent variable = "1" vs. "0". Standard errors are reported within the brackets.

Table 8

Summary of the replication results of Extension 2 (conceptual replication of Johnson et al. (2002) based on logistic regression analysis

Dradictor		Μ	lodel 1			N	Aodel 2	
Fieuciói	Estimate	Ζ	р	OR [95% CI]	Estimate	Ζ	р	OR [95% CI]
Intercept	0.72 (0.14)	5.11	<.001	2.05 [1.56, 2.71]	0.28 (0.16)	1.79	0.072	1.33 [0.97, 1.81]
Framing: Positive – Negative	-1.73 (0.14)	-12.06	<.001	0.18 [0.13, 0.23]	-1.82 (0.26)	-7.04	<.001	0.16 [0.09, 0.27]
Default: No-Default – Opt-In	-0.47 (0.17)	-2.67	0.007	0.62 [0.44, 0.88]	-0.39 (0.23)	1.68	0.091	1.48 [0.94, 2.32]
Default: Opt-Out – Opt-In	0.15 (0.17)	0.87	0.383	1.16 [0.83, 1.63]	-0.59 (0.24)	2.52	0.011	1.80 [1.14, 2.86]
Interaction terms:								
(Positive – Negative) × (No-Default–Opt-In)					0.18 (0.35)	0.52	0.604	1.20 [0.59, 2.41]
(Positive – Negative) × (Opt-Out –Opt-In)					0.07 (0.36)	0.21	0.83	1.07 [0.53, 2.17]
(Positive – Negative) × (No-Default–Opt-In) (Positive – Negative) × (Opt-Out –Opt-In)					0.18 (0.35) 0.07 (0.36)	0.52 0.21	0.604 0.83	1.20 [0.59, 2.41] 1.07 [0.53, 2.17]

Note. Estimates represent the odds of the dependent variable = "1" vs. "0"; N = 966. Standard errors are reported within the brackets.

Table 9

Summary and comparison of findings of the current replication study and the target studies

Part	Target effect	Original effect size	Replication effect size	Replication summary
Dert 1. Johnson and Caldstein (2002) *	Default effects: No-Default vs. Opt-In	4.72 [2.03 , 10.96]	1.38 [1.06 , 1.80]	Signal-inconsistent, smaller
Part 1: Johnson and Goldstein (2003)	Default effects: Opt-Out vs. Opt-In	5.93 [2.48 , 14.20]	1.67 [1.27 , 2.19]	Signal-inconsistent, smaller
	Default effects: No-Default vs. Opt-In	3.29 [1.28, 8.45]	0.66 [0.47, 0.94]	Signal-inconsistent, opposite
Part 2: Johnson et al. (2002)	Default effects: Opt-Out vs Opt-In	4.31 [1.62, 11.46]	0.75 [0.53, 1.05]	No signal-inconsistent
	Framing effects: Positive vs. Negative	1.86 [0.76, 4.57]	15.30 [10.23, 23.40]	Signal-inconsistent, stronger

Note. Replication summary based on the criteria by LeBel et al. (2019). (*) The effect size [Odds ratio] for this target study was calculated based on 2-sample test for equality of proportions.

Table 10

Summary of the findings of Johnson et al. (2002) across original, direct replication, and conceptual replication studies

Original	study's findings	Direct r	eplication findings	Concep	tual replication findings
Signal	Directionality	Signal	Directionality	Signal	Directionality
Yes	Consistent	No	Inconsistent - Opposite direction	No	Consistent
Yes	Consistent	No	Consistent	Yes	Consistent
No	Consistent	Yes	Consistent	Yes	Inconsistent - Opposite direction
	Original Signal Yes Yes No	Original study's findingsSignalDirectionalityYesConsistentYesConsistentNoConsistent	Original study's findingsDirect rSignalDirectionalitySignalYesConsistentNoYesConsistentNoNoConsistentYes	Original study's findingsDirect replication findingsSignalDirectionalitySignalDirectionalityYesConsistentNoInconsistent - Opposite direction NoYesConsistentNoConsistentNoConsistentYesConsistent	Original study's findingsDirect replication findingsConcepSignalDirectionalitySignalDirectionalitySignalYesConsistentNoInconsistent - Opposite direction YesNoNoConsistentYesConsistentYesNoConsistentYesConsistentYes

Note. Directionality dimension summarizes the directional consistency of results across Default effects and Framing effects; Predicted directionality of framing effects: participants' consent rates are higher in the positive frame than negative frame condition; Predicted directionality of default effects: consent rates are higher in 'Opt-Out' and 'No-Default' experimental condition than 'Opt-In' experimental condition. Signal, indicates support for the hypothesis using null hypothesis significance testing (p < .05)



Results of direct replication of Johnson et al. (2002). Percentage of participants who agreed to be notified about healthrelated surveys in the future. (A) Percentage of participants participating in the health survey by frame. (B) Percentage of participants participating in the health survey by default conditions. (C) Percentage of participants participating in the health survey by frame and conditions.



Figure 3

Results of Extension 1. Percentage of participants who consented to organ donation between permanent vs. temporary choice scenarios. Note.* p < .05, ** p < .01, *** p < .001

expected to find that participants in the positive framing condition consent at a higher rates than participants in the negative framing condition. While the original study did not find this, we found that consent rates were higher in positive frame condition than negative frame condition in our direct replication. However, in our conceptual replication, we found a framing effect in the opposite direction.

General discussion

We conducted a direct, close replication of Johnson and Goldstein (2003) and Johnson et al. (2002). In Part 1 of our study, we successfully replicated Johnson and Goldstein (2003). Participants consented to be organ donors at higher rates when they had to opt out of consent relative to participants who had to opt in. We found that participants in the No-Default condition—where no response was pre-selected—consented to organ donation at higher rates relative to participants who had to opt in. Additionally, we found that the permanence of these decisions affected people's choices.

Our replication results are consistent with Johnson and Goldstein (2003)—though the effects were smaller than those reported in the original study. The weaker effect is in line with recent work which found that effect sizes in large-scale studies were smaller than the estimates forecasted by academic experts and practitioners with relevant knowledge of nudge effects (DellaVigna



Extension 2: Percentage of participants who agreed to be notified about further information about organ donation in the future. (A) Participation rates by frame. (B) Participation rates by default conditions. (C) Participation rates by frame and conditions.

and Linos, 2022). Our well-powered study provides a more precise estimation of the effect size (OR = 1.67, 95% CI [1.27, 2.19]) that may be useful for future meta-analyses and for policy applications.

In Part 2 of our study, our replication results of Johnson et al. (2002) were inconsistent with the original findings. Unlike the original study, we found framing effects, yet we found no evidence for default effects. Consistent with the original study, we found that participants in the positive framing conditions consented to receive organ donation information at a higher rate than participants in the negative framing condition. However, in our conceptual replication of Johnson et al. (2002) that we report as Extension 2, participants in the positive framing condition information at a lower rate than participants in the negative framing condition information at a lower rate than participants in the negative framing condition.

Our results on default effects were inconsistent with the original findings in Johnson et al. (2002): we had no evidence for default effects overall. Nonetheless, we did find some indication of default effects when scenarios were framed positively. For instance, within positive framing conditions, participants in the No-Default condition and Opt-Out condition consented to receive health-related information at a higher rate than participants in the Opt-In condition. The pattern of results was in the opposite direction within the negative framing conditions: participants in the No-Default condition and Opt-Out condition consented to receive healthrelated information at a lower rate than participants in Opt-In condition. Interestingly, we found the consistent pattern across positive and negative frames in the conceptual replication: although these differences were not significant, participants in the No-Default condition and Opt-Out condition consented to receive organ donation related information at a higher rate than participants in Opt-In condition. As such, our results suggest that the stability of default effects can vary depending on the framing of the decision scenario.

There are several possible explanations for the inconsistent findings in our replications of Johnson et al. (2002). First, the failure to replicate the default effects may have been due to insufficient sample size in Johnson et al. (2002), which involved only 235 participants—about 39 participants for each experimental condition. This small sample may have led to false-positive results and inflated the effect size. Moreover, the smaller sample size in the original article may have resulted in the failure to detect the framing effects and the interaction that we found.

Second, the differences could be a result of changing preferences toward participating in online surveys. The original study was published in 2002, and the experimental scenario involved consenting to be notified about health-related surveys in the near future. People's preferences for taking part in online surveys may have changed in the last two decades. Therefore, the differences in the results could be informed by the change in

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Figure 5

Effect sizes in Johnson and Goldstein (2003), Johnson et al. (2002), and the current replication. Estimates and confidence intervals are plotted on a natural logarithmic scale.

peoples' preferences. Given the other successful replication in Part 1 of our study, we think this explanation is unlikely, yet we cannot rule out this possibility.

A third related explanation may be due to carry-over effects resulting from the order of the replications. The failed replication of Johnson et al. (2002) was in Part 2 and followed the unrelated organ-donation scenario in Part 1. We acknowledge that there is the slight possibility that somehow the order of execution affected the findings in Part 2. We consider this unlikely; the findings were not noise-they reflected a clear pattern of framing effects over default effect-so it would seem improbable that the slight manipulation in Part 1 triggered such a major shift from default effects to framing effects in Part 2. In our study design, we also took measures to mitigate carry-over effects. In Part 1, participants responded to the organ donation scenarios of Johnson and Goldstein (2003). The participants were assigned to three between-participants scenarios: Opt-Out, Opt-in, No-default. After completing Part 1 of the experiment, participants were randomly assigned to one of six between-participants conditions related to Johnson et al. (2002) in Part 2. So, we find it unlikely that a carry-over effect occurred in such a complex betweenparticipants design. Furthermore, Samples 1 and 2 had slightly different procedures. Despite these differences, we report similar results across the sample (see Table S11 in the supplementary materials). Therefore, this possibility of carry-over effects is unlikely.

Finally, the lack of support for the default effects in the negatively framed scenarios of Johnson et al. (2002) may have been due to the fact that double-negatively framed questions (i.e., negatively framed in the Opt-in scenario) are more confusing to participants than the other conditions. However, this possibility too seems to be an unlikely explanation for the lack of default effects. First off, the original study carried the same double-negatively framed questions yet found support for default effects. While we recognize that doublenegative questions may have been taxing to follow, the relatively consistent effects within negatively framed default conditions suggest otherwise. Across the three conditions with negatively-framed descriptions, the participation rates were similar: Opt-Out (28%), Opt-In (34%), and No-Default (25%). The similar participation rates across default conditions within the negative frame suggest that comprehending double negatively framed questions do not explain our pattern of findings.

There are also some potential explanations for other inconsistencies we found in our replications. Interestingly, the direction of framing effects in our conceptual replication of Johnson et al. (2002) was in the opposite direction of that found in our direct replication of Johnson et al. (2002). Although this result is inconsistent with the original study, it may not be entirely surprising; previous work suggests that framing effects may vary across task contexts. For example, work by Zhen and Yu (2016) show that framing effect vary between vignette-based vs. reward-based decision tasks. Furthermore, previous work also found that the direction of framing effects may differ based the relative attractiveness of the alternatives (Chandrashekar et al., 2021; Chen and Proctor, 2017; Wedell, 1997), or the degree to which decision may have personal relevance to participants (Krishnamurthy et al., 2001). Future work on framing effects may further investigate whether different task contexts modulate the direction of framing effects

At a more theoretical level, Wedell (1997) accentuation hypothesis perhaps best describes the pattern of current results about framing effects. Wedell (1997) argues that people have a higher need for justification in a positively framed choice than in a negatively framed choice. This higher need for justification highlights the differences between alternatives. On this account, when the overall attractiveness or benefits of participating in a health survey is high, people in the positively framed choice will choose to participate at a higher rate. Alternatively, when the overall attractiveness of participating in a health survey is lower, participants in the positively framed choice will choose to participate at a lower rate. Our results across direct and conceptual replication of Johnson et al. (2002) support this account. In the direct replication of Johnson et al. (2002) using a healthcare survey scenario, we find an overall high participation rate of 60.4% across conditions, and we found that participation rates were higher in the positive frame condition. In the conceptual replication of Johnson et al. (2002) that involved an organ donation scenario, we found an overall lower participation rate of 44.6% across experimental conditions, suggesting that the overall attractiveness of the option of consenting to receive additional information on organ donation is lower. Interestingly, we found that participation rates were lower in the positively framed condition. Our findings suggest that future work on the default effect may benefit from paying closer attention to the accentuation hypothesis.

Conclusion

Overall, our effort to replicate Johnson et al. (2002) contributes to the extant literature by testing the stability of default effects. Since the publication of Johnson et al. (2002), there has not been much interest in further studying framing effects (Positive vs. Negative frame) together with default effects. We believe that our findings indicate that this is a promising area for future research.

The current findings underline the importance of well-powered preregistered replications and extensions of notable findings in the judgment and decisionmaking literature. Our results suggest that the stability of default effects depends on the framing and context of the decision scenario and therefore hold valuable implications for the study of default effects. Although work on default effects has deservedly garnered attention from both scholars and public policy practitioners in the last two decades, our work suggests that we need a more refined and contextualized understanding of defaults' effectiveness.

We propose two main assertions. First, the effect size of default effects is likely smaller than those documented in original studies (DellaVigna and Linos, 2022. Therefore, we need well-powered samples to study default effects to achieve greater precision in our effect size estimates. Second, framing seems to influence the direction of default effects. Future work on default effects should be aware that people's decision frame can influence defaults' effectiveness. We hope the current replication opens up a range of theoretical and empirical work that can further future work on default effects.

Author Contact

Subramanya Prasad Chandrashekar, prasad.chandrashekar@ntnu.no, ORCID: 0000-0002-8599-9241

Correspondence about this article should be addressed to Gilad Feldman at gfeldman@hku.hk, ORCID: 0000-0003-2812-6599

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Author Contributions

Nadia Adelina, Shiyuan Zeng, Yan Ying Esther Chiu, and Grace Yat-Sum Leung analyzed the original articles, wrote the pre-registrations, designed the replications and the extensions, and conducted an initial analysis of the results and write-up of the first draft. Boley Cheng guided and assisted the replication effort. Subramanya Prasad Chandrashekar and Paul Henne verified and extended analyses, integrated the studies, and wrote the final manuscript for submission. Gilad Feldman led the replication efforts, supervised each step, conducted the pre-registrations, ran data collections, provided feedback throughout, and edited the final manuscript for submission.

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Default effect replications and extensions: Supplementary

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Disclosures

Data collection

Data collection was completed before analyzing the data.

Conditions reporting

We report all the conditions we collected.

Variables reporting

All variables collected for this study are reported and included in the provided data.

Author bios

Subramanya Prasad Chandrashekar is a research assistant professor with the Lee Shau Kee School of Business and Administration at the Open University of Hong Kong. His research focuses on social status, lay-beliefs, and judgment and decision-making.

Paul Henne is an Assistant Professor of Philosophy at Lake Forest College. He is affiliated with the neuroscience program at Lake Forest College. He works on experimental philosophy.

Nadia Adelina, Shiyuan Zeng, Yan Ying Esther Chiu, and Yat Sum Leung were students at the University of Hong Kong during the academic year 2019-2020.

Boley Cheng was a teaching assistant at the University of Hong Kong psychology department during the academic year 2019-2020.

Gilad Feldman is an assistant professor with the University of Hong Kong psychology department. His research focuses on judgment and decision-making.

Corresponding author

Gilad Feldman, Department of Psychology, University of Hong Kong, Hong Kong SAR; <u>gfeldman@hku.hk</u>

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Authorship declaration:

Table S0. Authors' contribution

In the Table below, we employ CRediT (Contributor Roles Taxonomy) to identify the contribution and roles played by the contributors in the current replication effort. Please refer to the URL (<u>https://www.casrai.org/credit.html</u>) on details and definitions of each of the roles listed below.

			Nadia Adelina,		
			Shiyuan Zeng, Yan Ving Esthor Chiu. Vot	Rolov	Cilad
Role	Prasad	Paul	Sum Leung	Cheng	Feldman
Conceptualization			X	8	X
Pre-registration			Х		Х
Data curation					Х
Formal analysis	Х	Х	Х		
Funding acquisition					Х
Investigation	Х	Х	Х		Х
Methodology			Х		Х
Pre-registration peer review /					
verification	Х		Х	Х	Х
Data analysis peer	v		v		
Project administration	Λ		Λ	v	v
				Λ	
Resources S a france as	V	V	V		Λ
Software	Х	Х	X		
Supervision				Х	Х
Validation	Х	Х			
Visualization	Х	Х			
Writing-original draft	Х	Х			
Writing-review and editing	Х	Х			Х

Exclusion criteria for the two replication studies

- 1. Subjects indicating a low proficiency of English (self-report < 5, on a 1-7 scale);
- Subjects who self-report not being serious about filling in the survey (self-report < 4, on a 1-5 scale);
- 3. Subjects who correctly guessed the hypothesis of this study in the funnelling section;
- 4. Have seen or done the survey before;
- 5. Subjects who failed to complete the survey. (duration = 0, leave question blank);
- 6. Not from the United States;

Project Process Outline

The current replication is part of the mass pre-registered replication project, with the primary aim of revisiting well-known research findings in the area of judgment and decision making (JDM) and examining the reproducibility and replicability of these findings. The current replication study followed the same project outline as noted below. For each of the replication projects, researchers completed full pre-registrations, data analysis, and APA style submission-ready reports. The authors independently reproduced the materials and designed the replication experiment, with a separate pre-registration document. The researchers then peer-reviewed one another to try and arrive at the best possible design. Then, the lead and corresponding authors reviewed the integrated work and the last corresponding author made final adjustments and conducted the pre-registration and data collection. The OSF page of the project contains the Qualtrics survey design used for data collection with pre-registration documents submitted by each of the researchers. In the manuscript, we followed the most conservative of the pre-registrations.





Verification of Analyses

Initial analyses were conducted by the independent researchers, who used JAMOVI (jamovi project, 2018) or R for data analyses. In preparing this manuscript, the lead and corresponding authors verified the analyses in R.

Power analyses

The rationale for reconstructing the original dataset and re-running logistic regression: authors of the original studies did not report full statistical results necessary for calculating effect size and power analysis. Hence, we had to re-conduct the test based on the data (consent rates/participation rates in different experimental conditions) available in the original paper. Given the study, designs involved binary outcome variables frequency tables and figures noted in the original papers allowed us to reconstruct accurate data collected during the original studies.

We note consistencies related to the power analysis details reported as part of the preregistrations across Group A and Group B (OSF: Group A - https://osf.io/mhwbe/, Group B https://osf.io/j4rpc/.). Below we report a rectified power analysis based on the original study's results. Please refer to Appendix A of the current supplementary document for explanations of the differences across Group's (A & B) analysis and the power analysis steps reported below.

Steps for power analysis (Johnson & Goldstein, 2003):

- 1. Conducted Binomial Logistic Regression in Jamovi based on the reconstructed data file ("OrigOrganDonation.csv" available on the OSF page of the project).
- 2. We calculated the Odds Ratios of the Binomial Logistic Regression results in Jamovi.
- 3. Used the odds ratio for the power analysis with GPower (Faul et al., 2007).

Supplementary screenshots

Generating Odds ratios of the Binomial	Binomial Logi	stic Reg	gressio	n					
Logistic	Model Fit Measures			0	usrall Mad	al Tost	-		
Regression	Model Deviance	e AIC	R ² McF		df	D	-		
results in Iamovi	1 180	186	0.114	23.0	2	<.001	-8		
		0000							
	Model Coefficients								
							95% Confid	ence Interv	al
	Predictor	Estimate	SE	Z	р	Odds ratio	Lower	Upper	
	Intercept ExpCondition:	-0.298	0.275	-1.08	0.278	0.742	0.433	1.27	
	Neutral – Optin	1.638	0.436	3.75	<.001	5.146	2.188	12.10	
	OptOut – Optin	1.780	0.445	4.00	<.001	5.930	2.477	14.20	
	Note. Estimates rep	resent the lo	g odds of "I	DV = 1" vs.	"DV = 0"				
Power analysis done by GPower 3.1 default effect (Opt-out vs. Opt-in) in Johnson & Goldstein, 2003	<i>G*Power out</i> , z tests - Log Options: Analysis: Input: Output:	put gistic Larg A pi Tail Odc Pr(Y α er Pow R ² o X di X pa Crit Tota Actu	regre ge sar riori: ((s) Is rati (= 1 X rr pro ver (1- other) stribu arm o ical z al san ual po	ssion nple : Comp = 1) H b -β err X ution	z-Te bute i 10 · prol	st, Der require	nidenk ed sam	co (20 ple si = = = = = = = = = = =	07) with var corr ze Two 5.93 0.5 0.05 0 Normal 0 1 1.9599640 36 0.9565547
	G*Power out	put							
Power analysis	z tests – Lo	qistic	reare	ssion					
by GPower 3.1	Options:	Lard	ie sar	nple	z-Te	st. Der	nidenk	(20	07) with var corr
default effect	Analysis:		riori: (Com	ute i	reauire	ed sam	ple si	ze
(No-default vs.	Input:	Tail	(s)					=	Two
Opt-in) in		Odd	ls rati	o				=	5.146
Johnson &		Pr(Y	′=1 X	=1)⊦	10			=	0.5
Goldstein, 2003		αει	r pro	b				=	0.05
		Pow	ver (1-	-β eri	· prol)		=	0.95
		R ² c	ther 2	x	•			=	0
		X di	stribu	ution				=	Normal
		Хра	arm µ					=	0
		Хра	arm σ					=	1
	Output:	Crit	ical z					=	1.9599640

Total sample	size		= 3	9	
Actual power	=	0.9557553			

Steps for power analysis (Johnson, Bellman & Lohse, 2002):

- 1. Conducted a Binomial Logistic Regression analysis in Jamovi based on the reconstructed data file ("OrigHealthSurvey.csv" available on the OSF page of the project).
- 2. Used the odds ratio from Jamovi output for the power analysis with GPower.

rating Odda							95% Confide	ence Interval
	Predictor	Estimate	SE	Z	р	Odds ratio	Lower	Upper
s of the	Intercept	-0.258	0.323	-0.798	0.425	0.773	0.410	1.46
mial	Framing:							
stic	Positive – Negative	0.621	0.459	1.354	0.176	1,860	0.757	4,57
	DefaultCondition:							
ession	Neutral – Opt-in	1.192	0.481	2.481	0.013	3.294	1.284	8.45
ts in Iamovi	Opt-out – Opt-in	1.462	0.499	2.931	0.003	4.314	1.623	11.46
	Framing * DefaultCondition:							
	(Positive – Negative) * (Neutral – Opt-in)	0.614	0.784	0.783	0.434	1.848	0.397	8.60
	(Positive – Negative) * (Opt-out – Opt-in)	0.373	0.795	0.468	0.639	1.451	0.305	6.90
	Note, estimates represent the log odds of 104 -	- 1 13. 01 -	0					
	wore, estimates represent the log odds of 104 -		U					
er analysis Power 3.1	z tests – Logistic regression Options: Large sample z Analysis: A priori: Compu	–Test, D ute requ) Demide ired sa	nko (20 ample si	07) wi ze	th var co	orr	

(opt-out vs. opt- in) in Johnson, Bellman & Lohse, 2002	Output:	Odds ratio Pr(Y=1 X=1) H0 α err prob Power (1-β err prob) R ² other X X distribution X parm μ X parm σ Critical z Total sample size Actual power		4.3137 0.5 0.05 0.95 0 Normal 0 1 1.9599640 44 0.9545112
Power analysis by GPower 3.1 default effect (No-default vs. opt-in) in Johnson, Bellman & Lohse, 2002	z tests – Log Options: Analysis: Input: Output:	gistic regression Large sample z-Test, Demidenk A priori: Compute required sam Tail(s) Odds ratio Pr(Y=1 X=1) H0 α err prob Power (1-β err prob) R ² other X X distribution X parm μ X parm σ Critical z Total sample size ual power = 0.9511642	o (20 ple s = = = = = = = = = = = =	007) with var corr ize Two 3.2941 0.5 0.05 0.95 0 Normal 0 1 1.9599640 56
Power analysis by GPower 3.1 framing effect (positive vs. negative) in Johnson, Bellman & Lohse, 2002	z tests - Log Options: Analysis: Input: Output: Act	gistic regression Large sample z-Test, Demidenk A priori: Compute required sam Tail(s) Odds ratio Pr(Y=1 X=1) H0 α err prob Power (1-β err prob) R ² other X X distribution X parm μ X parm σ Critical z Total sample size ual power = 0.9511280	o (20 ple s = = = = = = = = = = = =	007) with var corr ize Two 1.8603 0.5 0.05 0 Normal 0 1 1.9599640 156

Table S1. Summary of power analysis for the original paper:

Original Article	Effect	Conditions	Calculated Sample Size	Actual Power
Johnson &	Default Effect	Opt-out - Opt-in	36	0.95
(2003)	Default Effect	No-default - Opt-in	39	0.95
	Default Effect	Opt-out - Opt-in	44	0.95
Johnson, Bellman & Lohse (2002)	Default Effect	No-default - Opt-in	56	0.95
	Framing Effect	Negative - Positive	156	0.95
Total required sample size			~156 or more	

Notes: We choose a sufficiently larger sample size to ensure sufficient power of 95% to detect effects noted in the target studies and the effects of the proposed extension hypothesis that included six between-subjects conditions (conceptual replication of Johnson et al., 2002).

Power sensitivity analyses

Post-hoc power sensitivity analysis based on our intended goal of combined sample size of 2000 participants indicate that the final sample has 96.93% power (at $\alpha = .05$) detect an small-medium effect size (Odds Ratio = 1.50).

Power sensitivity analysis output using GPower 3.1:

z tests – Logi	istic regression		
Options:	Large sample z-Test, Demidenko (2007) with var corr		
Analysis:	Post hoc: Compute achieved powe	er	
Input:	Tail(s)	=	Two
	Odds ratio	=	1.50
	Pr(Y=1 X=1) H0	=	0.2
	α err prob	=	0.05
	Total sample size	=	2000
	R² other X	=	0
	X distribution	=	Binomial
	X parm π	=	0.5
Output:	Critical z	=	1.9599640
	Power (1–β err prob)	=	0.9693039



Sample comparison between the original studies and our two studies

Table S2. Sample differences and similarities between the original studies and our replication

 samples

	Johnson & Goldstein (2003)	Johnson, Bellman & Lohse (2002)	Sample 1: American MTurk workers	Sample 2: American MTurk workers
Sample size	161	235	954*	966
Geographic origin	International	International	US Americans	US Americans
Gender	Not reported	124 males, 153 females	459 males, 495 females	459 males, 514 females
Average age (years)	Not reported	35.4	38.7	39.2
Age range (years)	Not reported	Not reported	19-76	19-76
Medium (location)	Computer (online)	Computer (online)	Computer (online)	Computer (online)

Compensation	\$4.00	Not reported	Nominal payment	Nominal payment
Year	2003	2002	2019	2019

Differences and similarities between participant sample in the original study and replication.

Note: (*) 480 out of 954 participants in the Mturk sample were assigned to conditions that aimed to replicate the original findings. The remaining 474 participants were assigned to experimental conditions designed as part of the extension.

Materials and scales related to replication part

Type of study

Johnson & Goldstein, 2003: Between subjects design Johnson, Bellman & Lohse, 2002: Experimental Manipulations (Mixed design).

Experimental design of the original articles

Participants were asked to imagine that they have just moved to a new state and that they are filling out paperwork related to their move. The instruction that noted that they are "filling out the required online registration forms" upon their arrival at the new state, which was not present in Johnson & Goldstein (2003). This addition aimed to make the transition from the organ donor form (Part 1) to the health survey (Part 2) related forms more coherent.

Johnson & Goldstein, 2003

Three studies were reported in Johnson and Goldstein's (2003) work to evaluate the effect of the default on the agreement rate of organ donations. The focus of the current replication study is on the first experiment, which investigates the effect of three default conditions (i.e. opt-out, opt-in and no-default condition) on organ donation rate through the format of an online survey. Experiment 1 from the original research used a 3 (default options) between-subject design. The respondents of an online experiment were assigned to one of the three conditions with different default options. Participants were asked whether they would be donors based on one of the three questions with varying defaults. Table 7 shows the three experimental conditions.

Table S3. Experimental Design of Johnson & Goldstein, 2003 and our replication studies

Default Effect (Johnson & Goldstein, 2003)

Participants were told to complete an online survey, in the survey, they were told to assume that they have just moved to a new state and were asked whether they would be donors based on one of the three questions with varying defaults. Participants were randomly assigned to answer 1 out of 3 different default conditions of the question.

Independent	IV condition 1:	IV condition 2:	IV condition 3:
Variable:	Opt-out condition	Opt-in condition	No-default (no
Default			default) condition
(between subjects)	Participants answered	Participants answered	Participants answered
	the following multiple-	the following multiple-	the following multiple-
	choice question:	choice question:	choice question:
	Assume you moved to a	Assume you moved to a	Assume you moved to a
	new state in which the	new state in which the	new state, therefore,
	default is that you are	default is that you are	you need to select
	an organ donor, you are	not an organ donor, you	enrollment as an organ
	therefore by default	are therefore by default	donor. Please choose
	enrolled as an organ	not enrolled as an	your preferred organ
	donor. Please choose	organ donor. Please	donor status.
	your preferred organ	choose your preferred	
	donor status:	organ donor status:	Participants were given
			two options:
	Participants were given	Participants were given	• YES - Organ
	two options:	two options:	donor
	• YES - Organ	• YES - Organ	• NO - Not organ
	donor	donor	donor
	• NO - Not organ	• NO - Not organ	
	donor	donor	
	The option <i>"YES</i> -	The option "NO - Not	No options were set as
	Organ donor" was set	organ donor" was set as	the default option.
	as the default option.	the default option.	r · · ·
Dependent Variable:	Scoring the DV:	1	1

Donation agreement	Two choices were given in the multiple-choice questions, which were coded	
rate	with numerical values for further calculation:	
	• YES - Organ donor = 1	
	• NO - Not organ donor = 0	
	Value '1' indicated consent for organ donation, whereas, '0' indicated	
	disagreeing to donate the organs.	

Johnson, Bellman & Lohse, 2002

Experiment 2 of Johnson, Bellman & Lohse, 2002 employed a 2(framing) x 3(default) between-subject design, wherein each respondent of an online health survey was randomly assigned to answer one of six variations of the same question at the end of the survey they just completed. The dependent variable essentially asked whether the respondent would like to be contacted in the future for opportunities to participate in more health surveys. Table 1 details the six experimental conditions.

Table S4. . Experimental Design of Experiment 2 in Johnson, Bellman & Lohse, 2002 and our replication studies

Default and Framing Effect (Experiment 2 in Johnson, Bellman & Lohse, 2002)

Participants first completed a health survey online and then were told to answer a question that evaluates the effect of framing and default on participant's preferences on whether or not to receive notification for more information about health survey.

Participants were randomly assigned to 1 out of 6 different default and framing conditions and were required to confirm their choices accordingly. Experimental conditions varied in the structure of the question, i.e., the structure of the question (the DV) presented to the participants at the end of the health survey varied on framing and defaults.

Independent Variable:	IV condition 1:	IV condition 2:
Framing	Positive framing condition	Negative framing condition
(between subjects)		
	The following statement was randomly	The following statement was randomly

	assigned to the participants	:	assigned to the	assigned to the participants :	
	Notify me about more health surveys.		Do NOT notify me about more health surveys		
Independent Variable:	IV condition 1:	IV condition	2:	IV condition 3:	
Default	Opt-out condition	Opt-in condi	tion	No default condition	
(between subjects)	Default option: Yes	Default option	n: No	No default option was	
		· · ·		given.	
Dependent Variable	DV item variations in fram	ing:	DV item var	iations in default:	
	Participants were required to whether or not to receive ne about more health surveys.	to choose otification	Participants were required to choose whether or not to receive notification about more health surveys.		
	Two choices were given and were coded with numerical values for further calculation. The coded values were different for the two IV conditions.		Two choices were given and were coded with numerical values for further calculation. The coded values were different for the three IV conditions.		
	• $Y_{PS} = 1$).	• Yes	- 1	
	• $No = 0$		• <i>No</i>	= 0	
	Negative framing condition : • Yes = 0 • No = 1		Opt-in condition : • Yes = 0 • No = 1		
	The across both framing conditions the response to the DV was scored as "1" when the response meant participants agreed to receive more information about health survey, and the value of "0" meant declining such an offer.		No default co Yes No The across th response to t "1" when th agreed to rec	endition : = 1 = 0 nree defaults conditions the he DV was scored as e response meant participants reive more information about	
			health survey declining suc	y, and the value of "0" meant	

Materials and scales related to the extensions

Table S5.	Experimental design of Extension 1: Organ donation scenario (Permanent vs.
Temporary	

IV1: Choice permanence	IV1: Permanent	IV1: Temporary	
 (permanent vs. temporary) Between-subjects This is an extension IV2: Default Options (opt-in, opt-out vs no-default) between-subjects 	Refer to the scenarios of the originals study noted above.	Scenarios in the temporary mirrored that of the original study except for one additional information: The participant is told that organ donation status is renewable and can be changed every 3 years. "Please note that you will need to renew your option every 3 years."	
IV2: Opt-in	Dependent Variable		
Participants need to choose whether they want to confirm or change the status of NOT to be an organ donor	DV title: participation Specific DV item: participant's participation to become or not become an organ donor ("want to be an organ donor" vs "do not want to be an organ donor")		
	 Organ donor: if partici options. 	ipant selects either one of the following	
IV2: Opt-out Participants need to choose whether they want to confirm or change the	 "CONFIRM - I want to be an organ donor." "CHANGE - I want to be an organ donor." Not organ donor: if participant selects either one of the following options. "CONFIRM - I do not want to be an organ donor." 		

status of To be an organ donor	• "CHANGE - I do not want to be an organ donor."
IV2: No-default	
Participants were merely	
required to choose with No	
prior default and they can	
decide whether to be an	
organ donor or not.	

Note: The extension was part of the Sample 1 data collection

Extension 2: Conceptual replication of Johnson et al. (2002).

Extension 2 mirrored the Experiment 2 of Johnson, Bellman & Lohse, 2002 : a 2(framing) x 3(default) between-subject design, wherein each respondent answer one of six variations of the same questions. The dependent variable essentially asked whether the respondent would like to receive further information on organ donation.

Table S6.	Experimental	design	of Exter	nsion	2
-----------	--------------	--------	----------	-------	---

Extension 2: Default and Framing Effect:

They were manipulated by changing the format of the statement with default and framing conditions. Participants were randomly assigned to 1 out of 6 different versions of the statements, and were required to confirm their choices accordingly.

Independent Variable 1 (IV1): Default	IV1 condition 1: Opt-out condition	IV1 condition 2: Opt-in condition	IV1 condition 3: No default condition		
Independent Variable 2 (IV2): Framing	IV2 condition 1: Positive framing condition	IV2 condition 2: Negative framing condition			
Versions of questions formats (randomly assigned to	Version 1: Opt-in & Negative framing 1. Default condition: Opt-	Version 2: No default & Negative framing 1. Default condition:	Version 3: Opt-out & Negative framing 1. Default condition:		

participants in	in	No default	Opt-out		
Qualtrics survey)	- Default option = YES	- No default option was given	- Default option = NO		
	 2. Framing condition: Negative framing The following statement was randomly assigned to the participants : "Do NOT send me more information about organ donation" Yes No 	 2. Framing condition: Negative framing The following statement was randomly assigned to the participants : "Do NOT send me more information about organ donation" Yes No 	 2. Framing condition: Negative framing The following statement was randomly assigned to the participants : "Do NOT send me more information about organ donation" Yes No 		
	Version 4: Opt-out & Positive framing	Version 5: No default & Positive framing	Version 6: Opt-in & Positive framing		
	 Default condition: Optout Default option = YES 	 Default condition: No default No default option 	 Default condition: Opt-in Default option = 		
	 2. Framing condition: Positive framing The following statement was randomly assigned to the participants : "Send me more information about organ donation" Yes 	 was given 2. Framing condition: Positive framing The following statement was randomly assigned to the participants : "Send me more information about organ donation" 	 NO 2. Framing condition: Positive framing The following statement was randomly assigned to the participants : "Send me more information about organ donation" 		
	Νο	YesNo	YesNo		

Dependent Variable:	Participants would indicate their preference of receiving notifications for additional organ donation information under different default and framing conditions, and the effects of default and framing would be reflected by the participation rate.
Participation rate	
	Opt-out condition (IV1):
	• $Yes = 1$
	• $No = 0$
	Opt-in condition (IV1):
	• $Yes = 0$
	• $No = 1$
	No default condition (IV1):
	• $Yes = 1$
	• $No = 0$

Original articles' results

We reproduced the results of the original study to help us accurately pin-point the effect sizes for the current replication and to ascertain the degree of reproducibility.

The results of Study 1 of Johnson & Goldstein (2003)

Table S7. The results of Binomial Logistic Regression

						95% Confidence Inter	
Predictor	Estimate	SE	Z	р	Odds ratio	Lower	Upper
Intercept	-0.298	0.275	-1.08	0.278	0.742	0.433	1.27
Conditions:							
No-default vs. Opt-In	1.638	0.436	3.75	<.001	5.146	2.188	12.10
Opt-Out vs. Opt-In	1.780	0.445	4.00	<.001	5.930	2.477	14.20

Note. N =161; Estimates represent the log odds of "DV = 1" vs. "DV = 0";

The results of Study 2 of Johnson, Bellman & Lohse, 2002

Table S8. The results of Binomial Logistic Regression

						95% Confidence Interval	
Predictor	Estimate	SE	Z	р	Odds ratio	Lower	Upper
Intercept	-0.258	0.323	0.798	0.425	0.773	0.410	1.46
Framing:							
Positive vs. Negative	0.621	0.459	1.354	0.176	1.860	0.757	4.57
Default Condition:							
No-default vs. Opt-in	1.192	0.481	2.481	0.013	3.294	1.284	8.45
Opt-out – Opt-in	1.462	0.499	2.931	0.003	4.314	1.623	11.46
Framing * Default Condition:							
(Positive vs. Negative) * (No- default vs. Opt-in)	0.614	0.784	0.783	0.434	1.848	0.397	8.60
(Positive vs. Negative) * (Opt- out vs. Opt-in)	0.373	0.795	0.468	0.639	1.451	0.305	6.90

Note. N =235;Estimates represent the log odds of " DV = 1" vs. " DV = 0"

Additional Results of Replication

 Table S9.
 Descriptive table of the participation rates.

Doplication Study	Europeinson tol Conditions			Combined replication sample		
Replication Study	Experimental Condition	115	n	Participation rate		
Replication of	Opt-in de	fault	488	62.5%		
Experiment 1 from	Opt-out default			73.5%		
Johnson & Goldstein (2003)	No-default (no	o default)	482	69.7 %		
	Ont in default	Positive Framing	324	88.6%		
Replication of	Opt-in default	Negative Framing	324	33.6%		
Experiment 2 from	Ont out default	Positive Framing	321	93.1%		
Johnson, Bellman &	Opt-out default	Negative Framing	319	27.6%		
Lohse (2002)	No-default (no default)	Positive Framing	320	93.4%		
		Negative Framing	312	25.3%		

Note. N = 1920;

			Mt	urk Sample 1	Mturk Sample 2	
Replication Study	Experimental Condition	Experimental Conditions		Participation	n	Participatio
			11	rate	11	n rate
Replication of Experiment	Opt-in c	lefault	161	59.0%	327	64.2%
1 from Johnson &	Opt-out default		157	71.3%	320	74.6%
Goldstein (2003)	No-default (no default)		162	67.3%	319	70.9%
	Opt-in default	Positive Framing	160	84.4%	164	92.7%
Dauliastian of Evasiment		Negative Framing	161	30.4%	163	36.8%
2 from Johnson Ballmon	Ont out default	Positive Framing	162	88.9%	159	97.5%
& Lohse (2002)	Opt-out default	Negative Framing	159	22.6%	160	32.5%
	No default (no default)	Positive Framing	157	91.1%	163	95.7%
	no-default (no default)	Negative Framing	155	26.5%	157	24.2%

Table S10. Descriptive table of the participation rates presented by Sample 1 and Sample 2

Note. Sample 1 (N) = 954; Sample 2 (N) = 966;

Table S11.

Summary of the replication results: Logistic regression analysis conducted separately for Sample 1 and Sample 2

Target	Predictor	Estimate	SE	Z statistic	р	Odds ratio with 95% C.I.
Mturk sample 1						
	Intercept	0.36	0.16	2.27	0.023	1.44 [1.05, 1.98]
Part 1 (Johnson & Goldstein, 2003)	Default: No-default – Opt-in	0.36	0.23	1.54	0.124	1.43 [0.91, 2.26]
2003)	Default: Opt-out – Opt-in	0.55	0.24	2.30	0.022	1.73 [1.09, 2.77]
	Intercept	-0.83	0.17	-4.83	0.000	0.44 [0.31, 0.61]
	Default: No-default – Opt-in	-0.20	0.25	-0.78	0.433	0.82 [0.50, 1.34]
	Default: Opt-out – Opt-in	-0.40	0.26	-1.57	0.116	0.67 [0.41, 1.10]
Part 2 (Johnson, Bellman &	Framing: Positive – Negative	2.51	0.28	9.07	0.000	12.34 [7.17, 21.24]
Louse, 2002)	Framing * Defaults:					
	(Positive – Negative) * (No-default–Opt-in)	0.83	0.43	1.92	0.055	2.30 [0.98, 5.39]
	(Positive – Negative) * (Opt-out –Opt-in)	0.80	0.42	1.90	0.057	2.21 [0.98, 5.03]
Mturk sample 2						
	Intercept	0.58	0.12	5.07	<.001	1.79 [1.43, 2.26]
Part 1 (Johnson & Goldstein, 2003)	Default: No-default – Opt-in	0.31	0.17	1.82	0.068	1.36 [0.98, 1.90]
2003)	Default: Opt-out – Opt-in	0.49	0.17	2.85	0.004	1.64 [1.17, 2.30]
	Intercept	-0.54	0.16	-3.33	0.001	0.58 [0.42, 0.80]
	Default: No-default – Opt-in	-0.60	0.25	-2.43	0.015	0.55 [0.34, 0.89]
	Default: Opt-out – Opt-in	-0.19	0.23	-0.81	0.416	0.83 [0.52, 1.31]
Part 2 (Johnson, Bellman &	Framing: Positive – Negative	3.08	0.34	9.03	<.001	21.74 [11.15, 42.42]
Lonse, 2002)	Framing * Defaults:					
	(Positive – Negative) * (No-default– Opt-in)	1.17	0.55	2.13	0.033	3.21 [1.10, 9.39]
	(Positive – Negative) * (Opt-out –Opt-in)	1.31	0.63	2.07	0.039	3.70 [1.07, 12.81]

Note. Estimates represent the odds of dependent variable = "1" vs. "0"; N (Sample 1)= 480; N (Sample 2) = 966;

Notes on Johnson & Goldstein (2003) replication:

In sample 1, participants in the No-Default condition were not more likely to consent to organ donation (67.3%) than participants the Opt-In condition (59.0%) (b = 0.36, p = .124, OR = 1.43, 95% CI [0.91, 2.26]). In sample 2, participants in the No-Default condition were more likely to consent to organ donation (70.9%) than participants the Opt-In condition (64.2%) (b = 0.31, p = .068, OR = 1.36, 95% CI [0.98, 1.90]).

Notes on Johnson et al. (2002) replication:

In sample 1, participants in the No-Default did not consent to receive health-related information (59.0%) at a higher rate than participants the Opt-In condition (57.3%) (b = -0.20, p = .433, OR = 0.82, 95% CI [0.50, 1.34]). In sample 2, participants in the No-Default condition consented to receive health related information (60.6%) at a lower rate than participants the Opt-In condition (64.8%) (b = -0.60, p = .015, OR = 0.55, 95% CI [0.34, 0.89]).

Table S12.

Johnson et al.'s (2002) replication : Descriptive table of the participation rates split by framing

	Default Experimental	Posi	tive frame	Negative frame		
Replication Study	Conditions	ions n Part		n	Participation rate	
Replication of Experiment	Opt-in default	324	88.6%	324	33.6%	
1 from Johnson &	Opt-out default	321	93.1%	319	27.6%	
Goldstein (2003)	No-default (no default)	320	93.4%	312	25.3%	

Table S13.

Summary of the Johnson et al.'s (2002) replication results: Logistic regression analysis conducted separately for each frame

Target	Predictor	Estimate	SE	Z statistic	р	Odds ratio with 95% C.I.
By Positive Frame						
	Intercept	2.13	0.38	5.58	<.001	8.40 [3.98, 17.74]
Part 1 (Johnson & Goldstein, 2003)	Default: No-default – Opt-in	0.61	0.29	2.14	0.032	1.85 [1.05, 3.24]
	Default: Opt-out – Opt-in	0.58	0.28	2.04	0.025	1.79 [1.03, 3.11]
By Negative Frame						
Part 1 (Johnson & Goldstein, 2003)	Intercept	-0.68	0.12	-5.47	<.001	0.51 [0.40, 2.26]
	Default: No-default – Opt-in	-0.40	0.17	-2.29	0.022	0.67 [0.47, 0.94]
	Default: Opt-out – Opt-in	-0.29	0.17	-1.66	0.096	0.75 [0.54, 1.05]

Additional Results of Extension hypotheses

Table S14. Descriptive table for extension hypotheses.

			Mt	urk Sample 1
Extensions	Experimental Cor	n	Participation rate	
Mturk Sample 1: Organ donor study in the	Opt	155	63.2%	
temporary organ donor condition	Opt-	161	72.7%	
	No-defa	ult (no default)	158	74.7%
	Opt in default	Positive Framing	164	17.7%
	Opt-III default	Negative Framing	163	57.1%
Mturk Sample 2: Organ donor scenario adopted	Ont out default	Positive Framing	159	29.6%
with both framing and default effects.	Opt-out default	Negative Framing	160	70.6%
	No-default (no	Positive Framing	163	27.6%
	default)	Negative Framing	157	66.2%

Note. Sample 1 (N) = 954; Sample 2 (N) = 966;

Framework for evaluation of the replications

Table S15. Criteria for evaluation of replications by LeBel et al. (2018). A classification of relative methodological similarity of a replication study to an original study. "Same" ("different") indicates the design facet in question is the same (different) compared to an original study. IV = independent variable. DV = dependent variable. "Everything controllable" indicates design facets over which a researcher has control. Procedural details involve minor experimental particulars (e.g., task instruction wording, font, font size, etc.).

Target similarity	Highly similar				Highly dissimilar
Category	Direct replication				Conceptual replication
Design facet	Exact replication	Very close replication	Close replication	Far replication	Very far replication
Effect/ Hypothesis	Same/similar	Same/similar	Same/similar	Same/similar	Same/similar
IV operationalization	Same/similar	Same/similar	Same/similar	Different	Different
DV operationalization	Same/similar	Same/similar	Same/similar	Different	Different
IV stimuli	Same/similar	Same/similar	Different	Different	
DV stimuli	Same/similar	Same/similar	Different		
Procedural details	Same/similar	Different			
Physical setting	Same/similar	Different			
Contextual variables	Different				

Figure S2. Criteria for evaluation of replications by LeBel et al. (2019). A taxonomy for comparing replication effects to target article original findings.



Note: LeBel et al. (2019) suggested a replication evaluation using three factors: (a) whether a signal was detected (i.e., the confidence interval for the replication Effect size (ES) excludes one), (b) consistency of the replication ES with the original study's ES, and (c) precision of the replication's ES estimate (see Figure S# in the supplementary material).

References

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- Johnson, E. J., Bellman, S., & Lohse, G. L. (2002). Defaults, framing and privacy: Why opting in-opting out. Marketing Letters, 13(1), 5-15.

Johnson, E. J., & Goldstein, D. (2003). Do defaults save lives? Science, 302, 1338–1339.

Appendix A

Explanation of the power analysis reported as part of the Pre-registration across Group A and Group B.

Please note that the power analysis conducted based on the original studies is reported in Group A and Group B's supplementary material documents submitted as part of the pre-registration. (see OSF links: Group A = <u>https://osf.io/5e3r8</u>; Group B = <u>https://osf.io/3kdqb</u>)

Why are there some differences between the pre-registration reported as part of the supplementary from the ones of Pre-registration of Group A & B?

Essentially results change a little based on the reference level specified as part of the data analysis. In our supplementary document, we report results based on the default effect's reference level = **Opt-in**, as this reference level allows for easy evaluations of the arguments and results reported in the original studies. It does not affect us in a big way because our final sample size is much bigger than any of the power analyses conducted based on different combinations of reference levels.

Power analysis of Johnson & Goldstein, 2003:

The power analysis reported as part of the supplementary material is very similar to the one preregistrations across Group A and Group B. The numbers differ across these documents because of the variations in the reference levels specified as part of data analysis using JAMOVI (it's the same in R).

Power analysis related to effect size reported in supplymentary document:

We reproduce the results reported as part of the supplementary document. Reference level set were: (DV= 0; default condition = Optin)

Binomial Logistic Regression	()	Results								
Q Deper	dent Variable	Binomial Log	istic Re	gressio	n					
Covari	ates	Model Fit Measures								
		Model Deviance	AIC	R ² _{McF}	_					
		1 180	186	0.114						
	1									
Factor	;	Omnibus Likelihood R	atio Tests							
→ 8ª d	efault condition	Predictor	χ²	df	р					
		default condition	23.0	2	< .001					
					[3]					
		Model Coefficients - D	W							
✓ Model Builder		-						95% Confid	ence Interval	
default condition → Blocks	k1 x	Predictor	Estimate	SE	Z	р	Odds ratio	Lower	Upper	
→• de	ault condition	Intercept default condition:	-0.298	0.275	-1.08	0.278	0.742	0.433	1.27	
	dd New Block	Neutral – Optin	1.638	0.436	3.75	< .001	5.146	2.188	12.10	
		- Option - Optin	1.760	0.443	4.00	K .001	3,950	2/4/1	14.20	
		Note. Estimates repr	esent the log	odds of "Dv	/ = 1 VS. D	V = 0"				
✓ Reference Levels										
Variable	Reference Level	References								
🐣 DV	0 ~	References								
		 The jamovi p 	roject (2021). jamovi. (V	ersion 2.2)	[Compute	r Software]. R	etneved from	https://www.ja	novi.org.
😪 default condition	Optin 🗸									

Power analysis related to effect size reported in Group A Pre-registration document: (<u>https://osf.io/5e3r8;</u> Page 16)

We reproduce the results reported as part of Group A's pre-registration document. Reference level set were: (DV=0; default condition = OptOut)

Binomial Logistic Reg	pression (Results									
	Q Dependent Variable → ③ DV	•	Binomial Logis	stic Reç	gressio	n						
	Covariates		Model Fit Measures									
	\rightarrow		Model Deviance	AIC	R ² McF	-0						
			1 180	186	0.114							
						-1						
	Factors		Omnibus Likelihood Rati	io Tests								
	→ 😪 default condition		Predictor	χ ²	df	р						
			default condition	23.0	2	< .001						
						[3]						
		8										
✓ Model Builder			Model Coefficients - DV						199310 1994	n 1000 m ²⁰		
redictors	Blocks			2005.07					95% Confid	lence Interval		
default condition	→ Block 1	<	Predictor	Estimate	SE	Z	p	Odds ratio	Lower	Upper		
	→ default condition		Intercept default condition:	1.482	0.350	4.229	< .001	4.400	2.2144	8.743		
	+ Add New Block		Neutral – OptOut	-0.142	0.487	-0.291	0.771	0.868	0.3339	2.255		
			Optin – OptOut	-1.780	0.445	-3.996	< .001	0.169	0.0704	0.404		
			Note. Estimates represe	ent the log (odds of "DV	/ = 1" vs. "D\	/ = 0"					
✓ Reference Levels												
Variable	Reference Level		Poforoncoc									
🐣 DV	0	~	References									
🔗 default condition	OptOut	~	[1] The jamovi pro	ject (2021).	. jamovi. (V	ersion 2.2) [Computer	Software]. Ret	rieved from	https://www.jar	novi.org.	
			[2] R Core Team (2	021). <i>R: A I</i>	anguage a	and environn	ment for sta	tistical compu	ting. (Version	n 4.0) [Compute	er software]. Ret	rieved from
			project ora (R)	narkanac n	atriavarl fre	nm MRAN c	nanchot 20	21-04-01)				

Power analysis related to effect size reported in Group B Pre-registration document: (<u>https://osf.io/3kdqb</u>; Page 27)

We reproduce the results reported as part of Group B's pre-registration document. Reference level set were: (DV= 0; default condition = Optin). The analysis and results of Group B are the same as the supplementary material (because the reference levels were the same).

Binomial Logistic Regressio	n	(Resu	ilts								
C	C → Depender	t Variable 💧	Binc	mial Log	istic Reg	gressio	n					
	Covariates		Model	Fit Measures								
			Mod	el Deviance	AIC	R ² _{McF}						
			1	180	186	0.114	_					
	Factors	/	Omnib	us Likelihood Rr	tio Tests							
	→ 😪 defau	It condition	P	redictor	χ²	df	р					
			defau	t condition	23.0	2	< .001					
							[3]					
At Madal Builder			Model	Coefficients - D'	V							
redictor:	Placks									95% Confid	ence Interval	
default condition	→ Block 1	×		redictor	Estimate	SE	Z	р	Odds ratio	Lower	Upper	
	→ • default	condition	Interc	ept It condition:	-0.298	0.275	-1.08	0.278	0.742	0.433	1.27	
	+ Add	New Block	Ne	utral – Optin	1.638	0.436	3.75	< .001	5.146	2.188	12.10	
✓ Reference Levels			Note.	Estimates repre	sent the log	odds of "D\	/ = 1" vs. "D	V = 0"				
Variable		Reference Level										
			Refe	rences								
			[1]	The jamovi pr	oject (2021)	. jamovi. (V	ersion 2.2)	[Compute	r Software]. R	etrieved from	https://www.	amovi.org.
a default condition		Optin 🗸									10110	
			[2]	project.org. (F	2021). R: A L packages r	etrieved fr	om MRAN s	ment for st snapshot 2	tatistical comp 2021-04-01).	outing. (Versio	on 4.0) (Comp	iter software].

Power analysis of Johnson et al. 2002:

Power analysis reported as part of the supplementary material is very similar to the one preregistrations across Group A and Group B. The numbers differ across these documents because of the variations in the reference levels specified as part of data analysis using JAMOVI (or in R).

Power analysis related to effect size reported in supplementary document:

We reproduce the results reported as part of the supplementary document. Reference level set were: (DV= 0; default condition = Optin, Framing = negative)

PartID	Q	Dependent Variable		Model Fit M	deasures			Ove	rall Model	Test				
		Participation (DV)		Model	Deviance	AIC	R ² McF	X ²	df	p				
	→	Covariates		1	247	259	0.116	32.5	5	< .001				
				Model Coe	fficients - Parti	ipation (D	V)						95% Confid	ence Interva
		Factors			Predic	tor		Estimate	SE	Z	p	Odds ratio	Lower	Upper
	→	Rraming		Intercept				-0.258	0.323	-0.798	0.425	0.773	0.410	1.46
		o o o o da do	8:	positiv	e – negative			0.621	0.459	1.354	0.176	1.860	0.757	4.57
				Opt-or	ut – Opt-in			1.462	0.499	2.931	0.003	4.314	1.623	11.46
Model Ruilder				Neutra	al – Opt-in			1.192	0.481	2.481	0.013	3.294	1.284	8.45
Reference Levels				Framing : (positi	* Defaults: ve – negative)	k (Opt-out	t – Opt-in)	0.373	0.795	0.468	0.639	1.451	0.305	6.90
ariable		Reference Level		Alota Esti	ve – negative)	r (Neutral	- Opt-in)	U.014	0.784	0.783	0.434	1.848	0.397	8.00
Participation (DV)		0	~	NOTE: ESU	mates represei	it the log c	ous or Par	acipation (ov	/ = 1 vs. r	anticipation	00) = 0			
		-												

Power analysis reported in Group A Pre-registration document: (<u>https://osf.io/5e3r8</u>; Page 18/19)

We reproduce the results reported as part of Group A's pre-registration document. Reference level set were: (DV= 0; default condition = OptOut; Framing = positive)

PartID	O Dependent Va	riable	Model Fit N	leasures									
Partic	-> 🚷 Participa	ion (DV) 🔒					Over	all Model 1	est				
	Couprinter		Model	Deviance	AIC	R ² _{McF}	χ²	df	р				
	→		1	247	259	0.116	32.5	5	< .001				
			Model Coet	fficients - Parti	ipation (DV							0.501 (512) (513)	
	Factors			Predic	tor		Estimate	SE	Z	р	Odds ratio	Lower	Upper
	→ 🔗 Framing		Intercept				2.1972	0.527	4.1693	< .001	9.000	3.2037	25.283
	🔗 Defaults		Framing:										
			negativ	ve – positive			-0.9933	0.650	-1.5287	0.126	0.370	0.1036	1.323
			Defaults:				0.0202	0.746	0.0370	0.070	0.070	0.0054	4104
			Oct.in	- Opt-out			-0.0282	0.740	-0.0578	0.970	0.972	0.0474	4.194
			Framing	c Defaults:			-1.0345	0.015	-2.5012	0.005	0.100	0.0474	0.550
vlodel Builder			(ive = positive) :	: (Neutral –	Opt-out)	-0.2415	0.910	-0.2655	0.791	0.785	0.1321	4.670
lodel Builder			inegas.				0.2725	0.705	0.4684	0.639	1.451	0.3054	6.898
odel Builder ference Levels			(negati	ive – positive) :	k (Opt-in –	Opt-out)	0.5725	0.195	0.4004				
odel Builder Iference Levels	R	eference Level	(negati Note. Estir	ive – positive) : mates represer	k (Opt-in −) t the log od	Opt-out) ds of "Parti	cipation (DV)	= 1" vs. "Pi	articipation (D	OV) = 0"			
odel Builder ference Levels le ticipation (DV)	R	eference Level	(negat (negat Note. Estir	ive – positive) : mates represer	 (Opt-in −) t the log od 	Opt-out) ds of "Parti	icipation (DV)	= 1" vs. "Pi	articipation (D	0V) = 0"			
todel Builder eference Levels ole rticipation (DV) aming	R (eference Level 0 v positive v	Note. Estin	ive – positive) : mates represer tion Check	k (Opt-in –) t the log od	Opt-out) ds of "Parti	cipation (DV)	= 1" vs. "P.	articipation (E	OV) = 0"			
todel Builder eference Levels ble articipation (DV) aming efaults	R (eference Level 0 v positive v Opt-out v	Assumpt	ive – positive) : mates represer tion Check	k (Opt-in –) t the log od	Opt-out) ds of "Parti	icipation (DV)	= 1" vs. "Pi	articipation (E	2V) = 0"			

Power analysis related to effect size reported in Group B Pre-registration document: (<u>https://osf.io/3kdqb;</u> Page 19)

We reproduce the results reported as part of Group B's pre-registration document. The reference level set were: (DV= 0; default condition = Opt-out; Framing = positive). Please note the group B reconstructed the data of the original study and used different labels (data is identical) as presented below (Reference level guide: Opt-out = subscribed; Neutral = No default; Opt-in = unsubscribed).

Binomial Logistic Regressio	on		()	Binom	ial Logis	tic Reg	gression
		ependent Variable	*	Model Fit M	leasures		
Partition 1		Participation (DV)	8	Model	Deviance	AIC	R ² _{McF}
	\rightarrow	oranace.			247	259	0.116
				Model Coe	fficients - Parti	cipation (D	DV)
	F	actors			Predi	ctor	
	\rightarrow	😪 Framing		Intercept			
		🔗 Defaults		: Framing:			
				negati	ve – positive		
				Defaults:			
			<u></u>	Neutra	il – Opt-out		
Nedal Builder				Opt-in	- Opt-out		
/ Model Balidei				Framing :	E Defaults:		1 12 11 12
✓ Reference Levels				(negat (negat	ive – positive) ive – positive)	≭ (Neutral ≭ (Opt-in	I – Opt-out) – Opt-out)
Variable		Reference Level		Note. Esti	mates represer	nt the log o	odds of "Parti
🐣 Participation (DV)		0	~				
🔗 Framing		positive	~	Assump	tion Check	S	
🔗 Defaults		Opt-out	~	Colline	arity Statistics		

				Ove	rall Model	Test
Model	Deviance	AIC	R ² _{McF}	χ²	df	р
1	247	259	0.116	32.5	5	< .001
			10			

						95% Confide	ence Interva
Predictor	Estimate	SE	Z	р	Odds ratio	Lower	Upper
Intercept	2.1972	0.527	4.1693	< .001	9.000	3.2037	25.283
Framing:							
negative – positive	-0.9933	0.650	-1.5287	0.126	0.370	0.1036	1.323
Defaults:							
Neutral – Opt-out	-0.0282	0.746	-0.0378	0.970	0.972	0.2254	4.194
Opt-in - Opt-out	-1.8343	0.619	-2.9612	0.003	0.160	0.0474	0.538
Framing * Defaults:							
(negative - positive) * (Neutral - Opt-out)	-0.2415	0.910	-0.2655	0.791	0.785	0.1321	4.670
(negative - positive) * (Opt-in - Opt-out)	0.3725	0.795	0.4684	0.639	1.451	0.3054	6.898

Note. Estimates represent the log odds of "Participation (DV) = 1" vs. "Participation (DV) = 0"

Assumption Checks

Original Article	Effect	Conditions	Actual Power	Sample Size based on rectified power analysis	Group A	Group B	
Johnson &	Default Effect	Opt-out - Opt-in	0.95	36	90	101	
Goldstein (2003)	Default Effect	No-default - Opt-in	0.95	39 Not calculated		90	
	Default Effect	Opt-out - Opt-in	0.95	44	87	82	
Johnson, Bellman & Lohse (2002)	Default Effect	No-default - Opt-in	0.95	56	Not calculated	Not calculated	
	Framing Effect	Negative - Positive	0.95	156	144	232	
Total required sa	mple size			~156 or more	~144 or more	~ 232 or more	

Fable S16. Summary of powe	r analysis for the original p	paper across Group A and Group B:
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Notes: OSF links: Group A = <u>https://osf.io/5e3r8</u> ; Group B = <u>https://osf.io/3kdqb</u>